### STAT 260 - NONPARAMETRIC STATISTICS Ying Sun Statistics (STAT) Program Computer, Electrical and Mathematical Sciences & Engineering (CEMSE) Division King Abdullah University of Science and Technology (KAUST)

# HOMEWORK IV

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This question is about illustrating the problems with polynomial bases. First run

to simulate some apparently innocuous x, y data.

# <r code=""> ====================================</r>	======================== #
par(mar = c(4, 4, 2, 2) + .1) ; plot(x, y)	<pre># plotting</pre>
#  ====================================	======================== #



Figure 1: Some apparently innocuous x, y data.

(a)

Fit 5th and 10th order polynomials to the simulated data using, e.g.,  $lm(y \sim poly(x, 5))$ .

### Solution:

```
# <r code> ======= #
poly5 <- lm(y ~ poly(x, 5)) ; poly10 <- lm(y ~ poly(x, 10))
# </r code> ========= #
```

Plot the x, y data, and overlay the fitted polynomials. (Use the predict function to obtain predictions on a fine grid over the range of the x data: only predicting at the data fails to illustrate the polynomial behavior adequately).



Figure 2: Data with overlay of the fitted polynomials (5th order in red and 10th order in blue).

We see how the 10th order fitted polynomial chase the data and how he lost himself between the initial points of x because we don't have data in that interval.

(c)

One particularly simple basis for a cubic regression spline is  $b_2(x) = x$  and  $b_{j+2}(x) = |x - x_j^*|^3$  for j = 1, ..., q-2, where q is the basis dimension, and the  $x_j^*$  are knot locations. Use this basis to fit a rank 11 cubic regression spline to the x, y data (using lm and evenly spaced knots).

Solution:

(d)

Overlay the predicted curve according to the spline model, onto the existing x, y plot, and consider which basis you would rather use.



Figure 3: Data with overlay of the fitted polynomials (5th order in red and 10th order in blue) and the predicted cubic regression spline curve, in green.

The results with the 5th order polynomial and with the rank 11 cubic spline are quite similar, but with the cubic spline the fit looks more smooth.

## Problem 2

Show that the  $\beta$  minimizing  $\|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|^2 + \lambda \boldsymbol{\beta}^\top \mathbf{S}\boldsymbol{\beta}$  is given by  $\hat{\boldsymbol{\beta}} = (\mathbf{X}^\top \mathbf{X} + \lambda \mathbf{S})^{-1} \mathbf{X}^\top \mathbf{y}$ .

Solution:

$$\begin{split} \| \mathbf{y} - \mathbf{X}\boldsymbol{\beta} \|^2 + \lambda \boldsymbol{\beta}^{\mathsf{T}} \mathbf{S}\boldsymbol{\beta} \\ (\mathbf{y} - \mathbf{X}\boldsymbol{\beta})^{\mathsf{T}} (\mathbf{y} - \mathbf{X}\boldsymbol{\beta}) + \lambda \boldsymbol{\beta}^{\mathsf{T}} \mathbf{S}\boldsymbol{\beta} \\ \mathbf{y}^{\mathsf{T}} \mathbf{y} - 2\boldsymbol{\beta}^{\mathsf{T}} \mathbf{X}^{\mathsf{T}} \mathbf{y} + \boldsymbol{\beta}^{\mathsf{T}} \mathbf{X}^{\mathsf{T}} \mathbf{X}\boldsymbol{\beta} + \lambda \boldsymbol{\beta}^{\mathsf{T}} \mathbf{S}\boldsymbol{\beta} \\ \mathbf{y}^{\mathsf{T}} \mathbf{y} - 2\boldsymbol{\beta}^{\mathsf{T}} \mathbf{X}^{\mathsf{T}} \mathbf{y} + \boldsymbol{\beta}^{\mathsf{T}} (\mathbf{X}^{\mathsf{T}} \mathbf{X} + \lambda \mathbf{S}) \boldsymbol{\beta}, \end{split}$$

Taking the derivative with respect to  $\boldsymbol{\beta}$  and setting to zero:

$$-2\mathbf{X}^{\top}\mathbf{y} + 2(\mathbf{X}^{\top}\mathbf{X} + \lambda \mathbf{S})\hat{\boldsymbol{\beta}} = \mathbf{0}$$
$$(\mathbf{X}^{\top}\mathbf{X} + \lambda \mathbf{S})\hat{\boldsymbol{\beta}} = \mathbf{X}^{\top}\mathbf{y}$$
$$\hat{\boldsymbol{\beta}} = (\mathbf{X}^{\top}\mathbf{X} + \lambda \mathbf{S})^{-1}\mathbf{X}^{\top}\mathbf{y}.$$

### Problem 3

Let X be an  $n \times p$  model matrix, S a  $p \times p$  penalty matrix, and B any matrix such that  $\mathbf{B}^{\top}\mathbf{B} = \mathbf{S}$ . If  $\tilde{\mathbf{X}} = [\mathbf{X}^{\top}, \mathbf{B}^{\top}]^{\top}$  is an augmented model matrix, show that the sum of the first *n* elements on the leading diagonal of  $\tilde{\mathbf{X}}(\tilde{\mathbf{X}}^{\top}\tilde{\mathbf{X}})^{-1}\tilde{\mathbf{X}}^{\top}$  is tr{ $\{\mathbf{X}(\mathbf{X}^{\top}\mathbf{X} + \mathbf{S})^{-1}\mathbf{X}^{\top}\}$ .

Solution:

$$ilde{\mathbf{X}}^{ op} ilde{\mathbf{X}} = [\mathbf{X}^{ op} \mathbf{B}^{ op}] egin{bmatrix} \mathbf{X} \ \mathbf{B} \end{bmatrix} = \mathbf{X}^{ op} \mathbf{X} + \mathbf{B}^{ op} \mathbf{B} = \mathbf{X}^{ op} \mathbf{X} + \mathbf{S}$$

$$\begin{split} \tilde{\mathbf{X}} (\tilde{\mathbf{X}}^{\top} \tilde{\mathbf{X}})^{-1} \tilde{\mathbf{X}}^{\top} &= \begin{bmatrix} \mathbf{X} \\ \mathbf{B} \end{bmatrix} (\mathbf{X}^{\top} \mathbf{X} + \mathbf{S})^{-1} [\mathbf{X}^{\top} \mathbf{B}^{\top}] \\ &= \begin{bmatrix} \mathbf{X} (\mathbf{X}^{\top} \mathbf{X} + \mathbf{S})^{-1} \mathbf{X}^{\top} & \mathbf{X} (\mathbf{X}^{\top} \mathbf{X} + \mathbf{S})^{-1} \mathbf{B}^{\top} \\ \mathbf{B} (\mathbf{X}^{\top} \mathbf{X} + \mathbf{S})^{-1} \mathbf{X}^{\top} & \mathbf{B} (\mathbf{X}^{\top} \mathbf{X} + \mathbf{S})^{-1} \mathbf{B}^{\top} \end{bmatrix} \end{split}$$

The upper left  $n \times n$  submatrix of  $\tilde{\mathbf{X}}(\tilde{\mathbf{X}}^{\top}\tilde{\mathbf{X}})^{-1}\tilde{\mathbf{X}}^{\top}$  is  $\mathbf{X}(\mathbf{X}^{\top}\mathbf{X} + \mathbf{S})^{-1}\mathbf{X}^{\top}$ . Therefore, the sum of the first *n* elements on the leading diagonal is the tr{ $\mathbf{X}(\mathbf{X}^{\top}\mathbf{X} + \mathbf{S})^{-1}\mathbf{X}^{\top}$ }.

## Problem 4

Read Section 4.2.4 and Section 4.3 of the textbook. The additive model of section 4.3 can equally well be estimated as a mixed model.

(a)

Write a function which converts the model matrix and penalty returned by tf.XD into mixed model form. Hint: because of the constraints the penalty null space is of dimension 1 now, leading to a slight modification of  $D_+$ .

### Solution:

First taking the function tf.XD (and dependent functions) from the book (sections 4.2.1 and 4.3.1)

```
# producing constrained versions of X_{j} and D_{J}
tf.XD <- function(x, xk, cmx = NULL, m = 2) { \# get X and D subject to constraint
 nk = length(xk)
                                                   # number of knots x_{k}
 X = tf.X(x,xk)[, -nk]
                                                    # basis model matrix X
 D = diff(diag(nk), differences = m)[, -nk]  # square root penalty matrix D
 if (is.null(cmx)) cmx = colMeans(X) # values to subtract from the columns of X
 X = sweep(X, 2, cmx)
                                        # subtracting cmx from columns of X
 list(X = X, D = D, cmx = cmx)
                                                       # returning objects
}
  # taking a sequence of knots and an array of x values to produce a model matrix
                                           for a piecewise linear function
tf.X <- function(x, xj) { # tf basis matrix given data x and knot sequence x_{j}
 nk = length(xj); n = length(x)
                                                               # lengths
                                           # creating empthy model matrix X
 X <- matrix(NA, n, nk)
 for (j \text{ in } 1:nk) X[, j] = tf(x, xj, j)
                                                  # filling model matrix X
                                                 # returning model matrix X
 Х
}
                                     # defining the basis functions b_{j}(x)
tf <- function(x, xj, j) {</pre>
                                                       # tf: tent function
 dj = xj * 0 ; dj[j] = 1# generating j-th tf from set defined by knots x_{j}
 approx(xj, dj, x)$y
                                          # performing linear interpolation
}
```

Now, writing the function

```
# converting returned constrained model and penalty matrices into mixed model form
mmform <- function(x, xk = NULL, k = 10, sep = TRUE) {</pre>
                                # using default number of knots, k = 10
 if (is.null(xk))
                                             # if x_{k} is null,
  xk = seq(min(x), max(x), length = k)
                                             # build a grid of knots
 xd = tf.XD(x, xk)  # computing constrained versions of X_{j} and D_{j}
 Xa = ULAD(A, A, D)
D = rbind(0, xd$D); D[1, 1] = 1
                                              # doing modifications
 X = t(solve(t(D), t(xd_X)))
                                           # computing model matrix X
 if (sep) list(X = X[, 1, drop = FALSE], Z = X[, -1], xk = xk)
 else list(X = X, xk = xk)
}
```

Using your function from part (a) obtain the model matrices required to fit the two term additive tree model, and estimate it using lme. Because there are now two smooths, two pdIdent terms will be needed in the random list supplied to lme, which will involve two dummy grouping variables (which can just be differently named copies of the same variable).

Solution:

```
# <r code> =======
  # generating constrained versions of X_{j} and D_{J} matrices for the variables
x_h <- mmform(trees$Height) ; x_g <- mmform(trees$Girth)</pre>
               # putting together, building model matrix X with intercept column
X \leq cbind(1, x_h^X, x_g^X)
Z_h <- x_h; Z_g <- x_g
                                      # taking square root penalty matrices D
g1 <- g2 <- factor(rep(1, nrow(X)))
                                               # length of X, number of rows
library(nlme)
                                                          # loading library
Y <- trees$Volume
                                                           # response vector
                              # fitting the mixed model with positive definite
                                          matrices structure of class pdIdent
                              #
model <- lme(Y ~ X - 1, random = list(g1 = pdIdent(~ Z_h - 1),</pre>
                                  g2 = pdIdent(~ Z_g - 1))
           )
```

 $\square$ 

### (c)

Produce residual versus fitted volume and raw volume against fitted volume plots.



Figure 4: (a): fitted volume against residues; (b): fitted volume against raw volume.

In (a) we see the values spread between -4 and 4, what is good (but not so good); and in (b) we see that the values are quite similar, considering the sample size of 31.

### (d)

#### Produce plots of the two smooth effect estimates with partial residuals.





Figure 5: (a): height against s(height); (b): girth against s(girth).

Analyze some datasets using GAMs and GAMMs (when necessary), and with the bayesian framework.

"Which" bayesian framework?

- JAGS (R package rjags): Gibbs sampling;
- INLA (R package INLA): Integrated Nested Laplace Approximation.

To compare and to follow a reasoning starting with a more simple model, some non-bayesian models will be fitted using the R package mgcv.

Following the reasoning, some more simple models can also be fitted, as linear models and mixed linear models.

### Datasets

#### 1. Trade union data

Data on 534 U.S. workers with eleven variables (SemiPar::trade.union).

R summary output for the dataset:

years.educ	south female	years.experience	union.member
Min. : 2.00	0:378 0:289	Min. : 0.00	0:438
1st Qu.:12.00	1:156 1:245	1st Qu.: 8.00	1: 96
Median :12.00		Median :15.00	
Mean :13.02		Mean :17.82	
3rd Qu.:15.00		3rd Qu.:26.00	
Max. :18.00		Max. :55.00	
wage	age	race occupati	on sector married
wage Min. : 1.000	age Min. :18.00	race occupati 1: 67 1: 55	on sector married 0:411 0:184
wage Min. : 1.000 1st Qu.: 5.250	age Min. :18.00 1st Qu.:28.00	race occupati 1: 67 1: 55 2: 27 2: 38	on sector married 0:411 0:184 1: 99 1:350
wage Min. : 1.000 1st Qu.: 5.250 Median : 7.780	age Min. :18.00 1st Qu.:28.00 Median :35.00	race occupati 1: 67 1: 55 2: 27 2: 38 3:440 3: 97	on sector married 0:411 0:184 1: 99 1:350 2: 24
wage Min. : 1.000 1st Qu.: 5.250 Median : 7.780 Mean : 9.024	age Min. :18.00 1st Qu.:28.00 Median :35.00 Mean :36.83	race occupati 1: 67 1: 55 2: 27 2: 38 3:440 3: 97 4: 83	on sector married 0:411 0:184 1: 99 1:350 2: 24
wage Min. : 1.000 1st Qu.: 5.250 Median : 7.780 Mean : 9.024 3rd Qu.:11.250	age Min. :18.00 1st Qu.:28.00 Median :35.00 Mean :36.83 3rd Qu.:44.00	race occupati 1: 67 1: 55 2: 27 2: 38 3:440 3: 97 4: 83 5:105	on sector married 0:411 0:184 1: 99 1:350 2: 24

Colors by union.member status:

years.educ	0.35	0.38	0.15
	years.experience	0.087	0.98
		wage	0.18
	A REAL PROPERTY OF		age

Figure 6: Scatter plots and correlations between the numerical variables of the dataset trade.union.

### 2. Sitka spruce data

13 measurements of log-size for 79 Sitka spruce trees grown in normal or ozone-enriched environments. The first 54 trees have an ozone-enriched atmosphere, the remaining 25 trees have a normal (control) atmosphere. (SemiPar::sitka).



Figure 7: For illustration, Sitka spruce tree.

R summary output for the dataset:

id.num	order	days	log.size	ozone
Min. : 1	Min. : 1	Min. :152.0	Min. :2.230	0:325
1st Qu.:20	1st Qu.: 4	1st Qu.:227.0	1st Qu.:4.945	1:702
Median :40	Median : 7	Median :528.0	Median :5.630	
Mean :40	Mean : 7	Mean :441.8	Mean :5.547	
3rd Qu.:60	3rd Qu.:10	3rd Qu.:613.0	3rd Qu.:6.250	
Max. :79	Max. :13	Max. :674.0	Max. :7.560	

In the next par of graphs, each line correspond to a Sikta spruce tree along the evaluations, in days.

More comments are (will be) given in the project proposal presentation.

## Normal atmosphere

## Ozone-enriched atmosphere

